

## APPENDIX

### Formula 1 Prediction interval

In order to calculate the 95% prediction interval, the summary meta-analysis estimate  $M$ , the two sided critical t-value  $t_{1-0.05/2, k-1}$  and the standard deviation for the prediction interval  $SD_{PI}$  are needed. Here,  $t$  is the two-sided critical t-value that can be calculated via

<http://www.danielsoper.com/statcalc3/calc.aspx?id=10>. Fill in  $DF=k-1$  and probability level 0.025, with  $k$  the number of studies in the meta-analysis.  $SD_{PI}$  is the standard deviation of the prediction interval:  $SD_{PI} = \sqrt{(\tau^2 + SE^2)}$ , where  $\tau^2$  is the estimated heterogeneity and  $SE$  is the standard error of  $M$ <sup>1 18</sup>. If the  $SE$  was not reported, it can be approximated by dividing the distance between the limits of the 95% CI of the SMD by 3.92. The lower and upper limits of the 95% prediction interval are equal to  $M \pm t_{1-0.05/2, k-1} \times SD_{PI}$ . Of course it is possible to estimate prediction intervals with a different coverage, e.g. an 80% prediction interval would be based on  $t_{1-0.20/2, 6}$ . Note that the interval is calculated under the assumption that the value of  $\tau^2$  is known (and not estimated).

Estimations for ORs, risk ratios and hazard ratios are generally performed on the natural logarithm scale. As an example we take the calculation of a 95% prediction interval for an OR of 2.28 with a 95% CI from 1.05 to 4.96,  $\tau^2 = 0.353$  and  $k=7$ . The prediction interval will first be estimated on log scale. Note that the reported  $\tau^2$  is in general already the heterogeneity for log OR, not for OR, and can thus be used directly in the calculations. The  $SE$  of the log OR is calculated by dividing the distance between the log of the limits of the 95% CI of the OR by 3.92. This results in  $SE=0.318$ . The lower and upper limits of the 95% prediction interval for the log OR are  $\log(2.28) \pm 2.45\sqrt{(0.353 + 0.318^2)}$ . The value 2.45 results from the  $t_{1-0.05/2}$  distribution with 6 DF. Finally, we exponentiate the limits to return to the OR scale. The

resulting prediction interval ranges from 0.44 to 11.86, and can be interpreted as the 95% range of true ORs to be expected in similar studies.

### **Formula 2    Probability that effect is larger than threshold D**

The probability P that the true effect in a new study will be below a threshold D (e.g. the null effect) can be calculated with the left-tail cumulative t-distribution with k-1 degrees of freedom. The probability that the effect is above D equals  $1 - P$ .

In our example on nasal polyps the probability that the  $SMD \geq 0$  can be estimated as follows:

1. Start to calculate the probability P that a true  $SMD \leq 0$ . This is equivalent to the probability that a t-value  $\leq T$ , where T is equal to  $(D - M)/SD_{PI}$ , with summary treatment effect  $M = -0.51$ ,  $SD_{PI} = 0.425$  and  $D = 0$ . This results in  $T = 1.207$ , with 6 degrees of freedom (DF).
2. The probability P can be calculated online at <http://www.danielsoper.com/statcalc3/calc.aspx?id=41>. Fill in t value = 1.207 and DF = 6. The one-tailed probability  $P(t \leq 1.207) = 0.864$ .
3. We want the probability that the  $SMD \geq 0$ , this is  $1 - P = 0.136$ .

In the example on the OR (see formula 1), if we are interested in the probability of a null or negative effect, we are interested in the probability that a true  $OR \leq 1$ . For ORs, calculations must be based on the  $\ln OR$ , with  $M = \ln(2.28) = 0.824$ ,  $SD_{PI} = 0.674$ , and  $DF = 6$ . A true  $OR \leq 1$  corresponds to a true  $\ln OR \leq 0$ . Fill in  $T = (0 - 0.824)/0.674 = -1.223$  and  $DF = 6$ . The probability that a true  $OR \leq 1$  is equal to 0.134.

### **Formula 3    Prediction interval starting with $I^2$**

In order to calculate prediction intervals starting with an assumed  $I^2$  value (as percentage), we first calculated the corresponding  $\tau^2$  value:

$$\tau^2 = s^2 \frac{I^2}{100 - I^2}$$

with  $s^2$  the typical study variance, equal to  $\frac{\sum w_i (k-1)}{(\sum w_i)^2 - \sum w_i^2}$ , and  $w_i$  equal to the inverse of the study variance of study  $i$  ( $i=1..k$ ) and  $k$  the number of studies.<sup>23</sup> Subsequently formula 1 can be applied.

#### **Formula 4 Power of a future study**

Usually sample size calculations are performed without consideration of the heterogeneity. If we do take into account the heterogeneity, the expected power, i.e. the probability that a new study with  $N$  patients will have a positive result at significance level  $\alpha$ , given values for the standard error  $s$  of the new study and  $\mu$  and  $\tau^2$  as above, can be approximated with the delta method if  $\tau^2$  is not too large:

$$E(\text{power}) = g(\mu) + 0.5 \tau^2 g''(\mu)$$

where  $g$  is the power at the meta-analysis summary estimate  $\mu$ , and  $g''(\mu)$  is the second derivative of  $g$  at  $\mu$ . For  $g''(\mu)$  we can take the second derivative of the normal cumulative distribution function if  $N$  is sufficiently large.

This results in  $g''(\mu) = \frac{z_\mu e^{-0.5z_\mu^2}}{s^2 \sqrt{2\pi}}$ , with  $z_\mu = \frac{1.96s - \mu}{s}$ .

If the sample size  $N$  of the new study is such that the power for an effect of size  $\mu$  is 80%, the expected power of the study will be smaller than 80% if  $\tau^2$  is positive, because the corresponding value of  $z_\mu$  is negative.