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Mediterranean diet and risk of endometrial cancer: a pooled analysis of three italian case-control studies

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Background: Some components of the Mediterranean diet have favourable effects on endometrial cancer, and the Mediterranean diet as a whole has been shown to have a beneficial role on various neoplasms.

Methods: We analysed this issue pooling data from three case-control studies carried out between 1983 and 2006 in various Italian areas and in the Swiss Canton of Vaud. Cases were 1411 women with incident, histologically confirmed endometrial cancer, and controls were 3668 patients in hospital for acute diseases. We measured the adherence to the Mediterranean diet using a Mediterranean Diet Score (MDS), based on the nine dietary components characteristics of this diet, that is, high intake of vegetables, fruits/nuts, cereals, legumes, fish; low intake of dairy products and meat; high monounsaturated to saturated fatty acid ratio; and moderate alcohol intake. We estimated the odds ratios (OR) and the corresponding 95% confidence intervals (CI) for increasing levels of the MDS (varying from 0, no adherence, to 9, maximum adherence) using multiple logistic regression models, adjusted for major confounding factors.

Results: The adjusted OR for a 6–9 components of the MDS (high adherence) compared with 0–3 (low adherence) was 0.43 (95% CI 0.34–0.56). The OR for an increment of one component of MDS diet was 0.84 (95% CI 0.80–0.88). The association was consistent in strata of various covariates, although somewhat stronger in older women, in never oral contraceptive users and in hormone-replacement therapy users.

Conclusions: Our study provides evidence for a beneficial role of the Mediterranean diet on endometrial cancer risk, suggesting a favourable effect of a combination of foods rich in antioxidants, fibres, phytochemicals, and unsaturated fatty acids.

Although the most important determinant of endometrial cancer is elevated oestrogen levels (Liang and Shang, 2013), diet has been shown to have a role in the prevention of this neoplasm, possibly modifying oestrogen availability (Gaskins *et al*, 2009; World Cancer Research Fund and American Institute for Cancer Research, 2013).

The Mediterranean diet, typical of selected areas of the Mediterranean basin (including Italy, Southern France, Greece,

Spain and Morocco), is characterised by frequent consumption of various vegetables and fruit, cereals, fish and seafood, use of olive oil as the main seasoning fat, moderate alcohol (preferably in the form of wine) consumption, and relatively low consumption of meat and dairy products. Several epidemiological studies investigated the role of single food items characteristic of the Mediterranean diet on endometrial cancer risk, showing that high intake of vegetables and fruit (Bandera *et al*, 2007b), fibres and

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antioxidant vitamins (Goodman *et al*, 1997; Horn-Ross *et al*, 2003), cereals (La Vecchia *et al*, 1986; Potischman *et al*, 1993), legumes (Petridou *et al*, 2002) and fish (Brasky *et al*, 2014) has a favourable effect on the risk of this neoplasm, whereas high intakes of total, saturated and animal fat (Bandera *et al*, 2007a; Bravi *et al*, 2009), milk and dairy products (Ganmaa *et al*, 2012) increase the risk.

However, various components of the Mediterranean diet may act synergistically, leading to an overall effect greater than the sum of single foods, as found for other neoplasms, such as those of the oral cavity and pharynx (Bosetti *et al*, 2003; Filomeno *et al*, 2014), oesophagus (Bosetti *et al*, 2003), stomach (Praud *et al*, 2014), liver (Turati *et al*, 2014), pancreas (Bosetti *et al*, 2013) and larynx (Bosetti *et al*, 2003).

It has been estimated that about 10% of the incident endometrial cancers could be prevented if Western populations could shift their diet to the traditional Mediterranean diet (Trichopoulou *et al*, 2000). However, in a cohort of American women there was no relation between endometrial cancer risk and a recommended dietary score, based on the intake of 23 foods including a few components of Mediterranean diet such as fruit, vegetables, white meat, fish and whole grains (Mai *et al*, 2005). No reduction of endometrial cancer was found also in the Women's Health Initiative Dietary Modification Randomized Controlled Trial, whose goal was to reduce total fat intake and to increase intake of fruit, vegetables and grains (Prentice *et al*, 2007). An American case-control study considering four *a posteriori* defined dietary patterns (named 'plant-based', 'western', 'ethnic' and 'phytoestrogen rich') found no relation between any of these and endometrial cancer risk, although a diet characterised by a high fat consumption increased the risk (Dalvi *et al*, 2007). A Canadian case-control study found a weak inverse relation between endometrial cancer risk and a diet rich in plant foods and no relation with diets rich in sweets and meat (Biel *et al*, 2011).

Given the traditional high adherence to the Mediterranean diet of the Italian population, we investigated such issue pooling data from three case-control studies conducted in Italy and in the French speaking Switzerland, using a simple and intuitive *a priori* score of adherence to the Mediterranean diet proposed in the literature.

MATERIALS AND METHODS

Participants and study design. We analysed data from three case-control studies on endometrial cancer, conducted by the same group using similar protocols and questionnaires. Briefly, the first study was conducted in the greater Milan area and the provinces of Pordenone and Udine, Northern Italy, between 1982 and 1995 (La Vecchia *et al*, 1991), and included 538 cases of endometrial cancer and 2122 controls; the second study was conducted in the Swiss Canton of Vaud and in Northern Italy between 1988 and 1994 (Negri *et al*, 1996), and included 466 cases of endometrial cancer and 792 controls; and the third study was conducted in the greater Milan area and the provinces of Pordenone and Udine, in northern Italy, and the urban area of Naples, in southern Italy, between 1992 and 2006, and included 454 cases and 908 controls (Bravi *et al*, 2009). Cases were a total of 1458 women (median age 61 years) with incident, histologically confirmed endometrial cancer, admitted to the major teaching and general hospitals of the study areas; controls were 3822 women (median age 57 years) admitted to the same network of hospitals as cases for a wide spectrum of acute non-neoplastic conditions. Among controls, 29% were admitted for traumatic conditions, 29% for non-traumatic orthopaedic diseases, 18% for acute surgical conditions and 24% for other miscellaneous illnesses. Informed consent was obtained from each patient according to the local ethical rules. Among cases

and controls approached, <5% in Italy and 10% in Switzerland refused the interview. Since for 47 cases and 154 controls information was missing for at least one dietary item of interest (with the exception of legumes, for which the information was not available in the first study), our analyses were based on a total number of 1411 cases and 3668 controls.

Trained interviewers identified and questioned cases and controls in hospital, using structured questionnaire. These included information on socio-demographic and anthropometric characteristics, tobacco smoking, alcohol drinking, personal medical history, menstrual and reproductive factors, use of oral contraceptives and menopausal hormone-replacement therapy. Information on patients' usual diet in the 2 years before diagnosis was based on food frequency questionnaires, collecting information on weekly consumption of 14 selected indicator foods in the first study (La Vecchia *et al*, 1991), 50 foods in the second study (Negri *et al*, 1996) and 78 dietary items in the most recent study (Bravi *et al*, 2009). Intakes lower than once a week, but 1–3 times a month, were coded as 0.5 per week. The first two questionnaires have been shown to have a satisfactory reproducibility (D'Avanzo *et al*, 1997) and the most recent one a good reproducibility (Franceschi *et al*, 1995) and validity (Decarli *et al*, 1996) with the Spearman correlation coefficients between 0.60 and 0.80 for most food items considered in the Mediterranean Diet Score (MDS).

Dietary score. We used an *a priori* MDS based on nine characteristics of the traditional Mediterranean diet (Trichopoulou *et al*, 1995), that is, high consumption of fruit/nuts, vegetables, cereals (including bread and potatoes), legumes (for the second and third study only) and fish; high monounsaturated/saturated fatty acid ratio; low consumption of meat (including meat products) and milk (including dairy products); and moderate intake of alcohol (mainly wine). Subjective scores (low, medium and high) were used to obtain information on whole grain cereals in the first study and on monounsaturated and saturated fats in the first and second study. The study-specific median values according to the distribution of controls were used as cut-off points, for defining high and low intake of the nine components. A value of 1 or 0 was attributed to each subject for the presence of each characteristic. Thus, value 1 was given to: intake above median for fruit/nuts, vegetables, cereals, legumes, fish and high monounsaturated/saturated fatty acid ratio; intake below median for meat/meat products and milk/dairy products; and moderate consumption for alcohol, defined as intake over zero and below the study-specific median intake. The MDS was defined as the sum of the 9 individual binary components, and ranged between 0 (no adherence to the Mediterranean diet) and 9 (maximum adherence).

Statistical analysis. Odds ratios (OR) and corresponding 95% confidence intervals (CI) of endometrial cancer were computed for categories of the MDS using unconditional multiple logistic regression models (Breslow and Day, 1980). These included terms for age (quinquennia; categorically), study centre (Milan, Pordenone/Udine, Naples and Vaud), year of interview (continuous), education (<7, 7–<12, ≥12 years; categorically), body mass index (BMI, <21, 21–<25, 25–<30, ≥30 kg m⁻²; categorically), tobacco smoking (never, ex-smoker and current smoker of <14 or ≥14 cigarettes per day; categorically), total energy intake (quintiles; categorically), number of children (0, 1, >1; categorically), menopausal status/age at menopause (pre/perimenopausal, <50, ≥50 years; categorically), age at menarche (<12, 12–<13, >13 years; categorically), oral contraceptive use (never/ever), hormone-replacement therapy use (never/ever), history of hypertension (yes/no) and diabetes (yes/no). We also computed the OR for an increment of one component of the MDS. Analyses were conducted to investigate whether the association of the MDS with

endometrial cancer risk was heterogeneous across strata of selected covariates.

All the analyses were performed using the SAS software, version 9.2 (SAS Institute, Inc., Cary, NC, USA).

RESULTS

The distribution of 1411 endometrial cancer cases and 3668 controls according to age and other potential confounders is reported in Table 1. Cases were older, had a higher BMI, reported more frequently a history of diabetes and hypertension, had a

younger age at menarche and an older age at menopause, were more frequently never smokers, and never users of oral contraceptives and ever users of hormone-replacement therapy.

Table 2 shows the adjusted ORs of endometrial cancer risk for each component of the MDS. High vs low consumption of vegetables (OR 0.64), fruit/nuts (0.85) and monounsaturated to saturated fatty acid ratio (0.73), low vs high consumption of meat/meat (0.77) and moderate vs no/high alcohol consumption (0.76) were associated to a significant reduced risk of endometrial cancer, while high consumption of legumes (1.26) and low consumption of milk and dairy products (1.19) were associated to a significant increased risk. No significant association was found for high vs low consumption of cereals and potatoes (OR 0.99) or fish (1.00).

The distribution of endometrial cancer cases and controls and the corresponding ORs according to the MDS are given in Table 3. Compared with a 0–3 components of the MDS, the OR decreased for increasing levels of the MDS, with an OR of 0.43 (95% CI 0.34–0.56) for 7–9 MDS with a significant trend in risk

Table 1. Distribution of 1411 cases of endometrial cancer and 3668 controls according to age, education and other selected variables—Italy and Switzerland, 1983–2006

	Cases		Controls		P-value ^a
	N	%	N	%	
Age (years)					
<50	170	12.0	1011	27.6	<0.0001
50–54	179	12.7	506	13.8	
55–59	260	18.4	554	15.1	
60–64	290	20.6	590	16.1	
65–69	250	17.7	507	13.8	
≥70	262	18.6	500	13.6	
Education (years) ^b					
<7	758	53.8	1941	53.0	0.39
7–<12	424	30.1	1092	29.8	
≥12	226	16.1	631	17.2	
Body mass index (kg m ⁻²) ^b					
<25	526	37.5	2052	64.1	<0.0001
25–<30	478	34.0	1186	32.4	
≥30	401	28.5	422	11.5	
Smoking status ^b					
Never smokers	1075	76.2	2542	69.6	<0.0001
Ex-smokers	127	9.0	339	9.3	
Current smokers	208	14.8	769	21.1	
History of diabetes					
No	1216	86.2	3505	95.6	<0.0001
Yes	195	13.8	163	4.4	
History of hypertension					
No	760	53.9	2435	66.4	<0.0001
Yes	651	46.1	1233	33.6	
Age at menarche (years) ^b					
<12	240	17.0	593	16.2	0.027
12–13	641	45.6	1547	42.3	
>13	526	37.4	1518	41.5	
Number of children					
0	274	19.4	685	18.7	0.49
1	309	21.9	795	21.7	
>1	828	58.7	2188	59.6	
Age at menopause (years) ^b					
Pre/peri	259	18.4	1099	30.0	<0.0001
<50	374	26.6	1145	31.3	
≥50	772	55.0	1416	38.7	
Oral contraceptive use					
Never	1308	92.7	3261	88.9	<0.0001
Ever	103	7.3	407	11.1	
Hormone-replacement therapy use					
Never	1270	90.0	3491	95.2	<0.0001
Ever	141	10.0	177	4.8	

^aP-value for association for dichotomous variables and P-value for trends for variables with more than two categories from χ^2 statistic.

^bThe sum does not add up to the total because of some missing values.

^aP-value for association for dichotomous variables and P-value for trends for variables with more than two categories from χ^2 statistic.

^bThe sum does not add up to the total because of some missing values.

Table 2. ORs and 95% CIs for endometrial cancer among 1411 cases and 3668 controls according to the nine dietary items included in the MDS—Italy and Switzerland, 1983–2006

MDS components	Comparison level ^a	OR (95% CI) ^b
Vegetables	High vs low	0.64 (0.55–0.75)
Fruit/nuts	High vs low	0.85 (0.73–0.98)
Legumes	High vs low	1.26 (1.00–1.58)
Cereals and potatoes	High vs low	0.99 (0.82–1.20)
Fish	High vs low	1.00 (0.86–1.17)
Monounsaturated to saturated fatty acid ratio	High vs low	0.73 (0.63–0.84)
Meat and meat products	Low vs high	0.77 (0.67–0.88)
Milk and dairy products	Low vs high	1.19 (1.04–1.36)
Alcoholic beverages	Moderate vs no/high	0.76 (0.65–0.87)

Abbreviations: CI = confidence interval; MDS = Mediterranean diet score; OR = odds ratio.
^aHigh: consumption above the study-specific median; low: consumption below the study-specific median; moderate: consumption higher than zero but below the study-specific median.

^bEstimated from logistic regression models adjusted for age, study centre, year of interview, education, tobacco smoking, body mass index, age at menopause, age at menarche, parity, oral contraceptive use, hormone-replacement therapy use, history of hypertension, diabetes and total energy intake.

Table 3. ORs and 95% CI for endometrial cancer among 1411 cases and 3668 controls according to the MDS—Italy and Switzerland, 1983–2006

No. of MDS components	Cases	Controls	OR (95% CI) ^a
0–3	311	683	1.00 ^b
4	329	857	0.82 (0.67–1.00)
5	343	929	0.66 (0.54–0.81)
6	250	677	0.54 (0.43–0.68)
7–9	178	522	0.43 (0.34–0.56)
P-value for trend	—	—	<0.0001
OR ^c	—	—	0.84 (0.80–0.88)

Abbreviations: CI = confidence interval; MDS = Mediterranean diet score; OR = odds ratio.

^aEstimated from logistic regression model adjusted for age, study centre, year of interview, education, tobacco smoking, body mass index, age at menopause, age at menarche, parity, oral contraceptive use, hormone-replacement therapy use, history of hypertension, diabetes and total energy intake.

^bReference category.

^cOR for one unit increment of the MDS.

(P for trend <0.0001). The OR for an increment of one unit of the MDS, that is, one component of the Mediterranean diet, was 0.84 (95% CI 0.80–0.88).

The association between the increment of one unit of the MDS and endometrial cancer risk were analysed in strata of selected covariates (Table 4). Risk estimates were consistent across strata of education, BMI, tobacco smoking, history of diabetes and hypertension, age at menarche, parity and age at menopause, while the association was significantly stronger in the second study (OR 0.76, 95% CI 0.69–0.84), in women aged ≥ 60 years (0.80, 0.75–0.85), in never oral contraceptive users (0.83, 0.78–0.87) and in ever hormone-replacement therapy users (OR 0.70, 0.58–0.83).

DISCUSSION

In the present study, based on a large case-control investigation, we found a reduced risk of endometrial cancer for increasing adherence to the Mediterranean diet, with an over 50% risk reduction for women in the highest vs the lowest score. Considering the single components of the Mediterranean diet, we found significant inverse associations only for high consumption of vegetables and fruit, high ratio of monounsaturated to saturated fatty acids, low intake of meat and moderate alcohol consumption. For milk and legumes, we found associations even opposite to those expected. This clearly supports the hypothesis that Mediterranean diet as a whole is a stronger determinant of endometrial cancer risk than the single dietary components, probably because dietary score take into account the interactions among various combinations of foods and nutrients and their synergistic effects.

Our results show a greater favourable effect of the Mediterranean diet on endometrial cancer risk than that found in previous investigations (Mai *et al*, 2005; Dalvi *et al*, 2007; Prentice *et al*, 2007). This possibly depends, at least in part, on the definition of the MDS (Trichopoulou *et al*, 1995). In fact, the MDS is defined as a score in relation to the median consumption of selected foods in the population of controls and not in absolute terms (i.e., amount of foods in grams). However, subjects with a higher MDS score in this Italian population are likely to have higher absolute intake of Mediterranean food components as compared with other populations, as in Italy consumption of the various components of the Mediterranean diet is traditionally higher.

The Mediterranean diet is rich in vitamins, carotenoids, flavonoids and folates (mainly derived from vegetables and fruit), which have shown inverse relations with endometrial cancer in most case-control studies, although no relation in prospective ones (Jain *et al*, 2000; Pelucchi *et al*, 2008; Bandera *et al*, 2009; Bidoli *et al*, 2010; Cui *et al*, 2011; Aarestrup *et al*, 2012; Harris *et al*, 2012; Tavani *et al*, 2012). These compounds are rich in antioxidants and have anticarcinogenic properties. Moreover, foods characteristic of the Mediterranean diet are also rich in fibres, which have been involved in oestrogen metabolism (Xu *et al*, 2007). High intake of fibres decreases serum oestrogen levels (Rose, 1990) and vegetarian women have lower plasma oestrogen levels and lower urinary excretion of estrogens than omnivores (Goldin *et al*, 1982; Gorbach and Goldin, 1987), thus lowering the risk of endometrial cancer (Key and Pike, 1988). Moreover, a large proportion of energy intake in the Mediterranean diet derives from cereals and other plant sources, which has been inversely related with endometrial cancer risk rather than from animal sources, with a high content of saturated fats (Xu *et al*, 2007). Dietary fats (mainly saturated from animal sources) may influence the metabolism of oestrogens by increasing the intestinal re-absorption of oestrogens (Gorbach and Goldin, 1987) (and consequently their serum levels) and, thus, the

Table 4. OR and 95% CI of endometrial cancer among 1411 cases and 3668 controls according to the MDS in strata of selected covariates—Italy and Switzerland, 1983–2006

	Cases/ Controls	OR (95% CI) ^a	P-value for interaction ^b
Age (years)			
< 60	609/2071	0.89 (0.83–0.96)	0.011
≥60	802/1597	0.80 (0.75–0.85)	
Study			
First study	524/1194	0.83 (0.76–0.90)	0.0048
Second study	436/772	0.76 (0.69–0.84)	
Third study	451/902	0.94 (0.86–1.02)	
Education (years) ^c			
<7	758/1941	0.83 (0.77–0.88)	0.45
≥7	650/1723	0.85 (0.80–0.91)	
Body mass index (kg m ⁻²) ^b			
< 25	526/2052	0.86 (0.81–0.92)	0.20
≥25	879/1608	0.82 (0.77–0.87)	
Tobacco smoking ^b			
Never smoker	1075/2542	0.85 (0.80–0.90)	0.18
Ex-smoker	127/339	0.89 (0.76–1.04)	
Current smoker	208/769	0.77 (0.69–0.86)	
History of diabetes			
No	1216/3505	0.85 (0.80–0.89)	0.35
Yes	195/163	0.78 (0.67–0.91)	
History of hypertension			
No	760/2435	0.84 (0.79–0.90)	0.84
Yes	651/1233	0.83 (0.78–0.90)	
Age at menarche (years) ^c			
< 12	240/593	0.83 (0.74–0.93)	0.83
≥12	1167/3065	0.84 (0.80–0.89)	
Parity			
Nulliparae	274/685	0.81 (0.74–0.90)	0.47
Parae	1137/2983	0.85 (0.80–0.89)	
Age at menopause (years) ^c			
Pre/peri menopause	259/1099	0.89 (0.81–0.98)	0.20
< 50	374/1145	0.86 (0.79–0.94)	
≥50	772/1416	0.81 (0.76–0.86)	
Oral contraceptive use			
Never	1308/3261	0.83 (0.78–0.87)	0.015
Ever	103/407	1.01 (0.86–1.17)	
Hormone replacement therapy			
Never	1270/3491	0.85 (0.81–0.90)	0.025
Ever	141/177	0.70 (0.58–0.83)	

Abbreviations: CI = confidence interval; MDS = Mediterranean diet score; OR = odds ratio.

^aOR for one unit increment of the MDS, estimated from logistic regression model adjusted for age, study centre, year of interview, education, tobacco smoking, body mass index, age at menopause, age at menarche, parity, oral contraceptive use, hormone replacement therapy use, history of hypertension, diabetes and total energy intake.

^bP-value for heterogeneity tested by likelihood ratio test and corresponding χ^2 statistic.

^cThe sum does not add up to the total because of some missing values.

Abbreviations: CI = confidence interval; MDS = Mediterranean diet score; OR = odds ratio.

^aOR for one unit increment of the MDS, estimated from logistic regression model adjusted for age, study centre, year of interview, education, tobacco smoking, body mass index, age at menopause, age at menarche, parity, oral contraceptive use, hormone replacement therapy use, history of hypertension, diabetes and total energy intake.

^bP-value for heterogeneity tested by likelihood ratio test and corresponding χ^2 statistic.

^cThe sum does not add up to the total because of some missing values.

risk of endometrial cancer (Key and Pike, 1988). This is in agreement with the protective effect found in our study for a high unsaturated to saturated fat ratio. The direct association of red meat intake with cancer risk has been attributed to heterocyclic amines, polycyclic aromatic hydrocarbons and nitrosamines produced during cooking, though the interpretation remains open to discussion (Tavani *et al*, 2000).

The effect of Mediterranean diet on endometrial cancer seems to be independent from its eventual effect on body weight, as Mediterranean diet has not been related with BMI and waist-to-hip ratio (Rossi *et al*, 2008). However, we carefully allowed for BMI in our analyses.

Limitations of this study are those of all case-control studies, including selection and information bias (Breslow and Day, 1980). However, the participation rates were high and similar between cases and controls and the catchment area was also comparable. The study was hospital-based, but we excluded from controls all subjects with admission diagnosis related to known risk factors for endometrial cancer and hormonal disorders, and information from hospital patients has shown satisfactory reproducibility (D'Avanzo *et al*, 1997). Dietary habits of hospital controls may not be representative of the general population, but we have also excluded women with chronic conditions which may have altered diet or body weight. We used two food frequency questionnaires tested for reproducibility (Franceschi *et al*, 1993; D'Avanzo *et al*, 1997) and validity (Decarli *et al*, 1996) and there is no reason for a different recall between cases and controls, as both were interviewed in the same hospital setting. We used an *a priori* developed and commonly used MDS (Trichopoulou *et al*, 2003). Our risk estimates were allowed for many confounding variables, including endogenous and exogenous hormonal factors and body weight, and the consistency of the direct association through most strata of covariates excludes a major role of residual confounding. The strength of the inverse association, the observed linear relation between risk and exposure, and the biological plausibility support a causality of the relation, although the evidence is still scanty and needs to be confirmed in other studies mainly prospective ones.

In conclusion, our data suggest that a Mediterranean dietary pattern, more than single foods, may favourably influence the risk of endometrial cancer. This supports the hypothesis that a diet rich in a combination of vegetables, fruit and cereals, poor in saturated fats, meat and dairy product, and with a moderate intake of wine has a favourable role on endometrial cancer risk.

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

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